

## PROBABILITY: QUICK REVIEW

### TWO conceptions:

- certainty/likelihood of particular events  
e.g.,  $\text{Pr}(\text{particular coin flip} = \text{Heads})$  ;  
e.g.,  $\text{Pr}(i^{\text{th}} \text{ word in a certain text is a noun})$
- relative frequency of a certain kind of outcome in an infinitely large number of repetitions of some "repeatable experiment"  
e.g., proportion of "heads" in  $\infty$  no. of coin flips  
e.g., proportion of nouns in an  $\infty$  text corpus
- In estimating probabilities, we use the latter conception.
- In applying (estimated) probabilities, we use the first conception.

## RANDOM VARIABLES

- A R.V. picks out some observable in a repeatable experiment
- Its value varies from repetition to repetition  
e.g., "Flip" could be = H (for heads), = T (for tails) in repeated coin tosses  
e.g., "CAT" could be N, V, P, A, Det, ... (i.e., P.O.S.) for a word found in some arbitrary location in a corpus

$$\text{Pr}(\text{Flip} = \text{H}) = .497$$

says that the prob. of heads in flipping some coin is 49.7%

$$\text{Pr}(\text{CAT} = \text{N}) = .24$$

says that the prob. that an arbitrarily selected word in some large corpus (or in some type of corpus) is a noun is 24%.

Abbreviation :

$$\text{Pr}(\text{H}) = .497, \quad \text{Pr}(\text{N}) = .24$$

r.v. is "understood"

## Jointly distributed r.v.'s

- We can consider multiple r.v.'s at the same time, coming from one experiment or multiple experiments
- The values they assume may be dependent on one another (because there is some causal connection between the r.v.'s), or independent. (We'll formalize this.)

E.g., Let  $(X_1, X_2)$  be r.v.'s for the outcomes of 2 successive rolls of a die.

So, values are  $(1, 1), (1, 2), \dots, (6, 6)$

For a fair die & causally independent, randomized rolls,  $\Pr(X_1=1, X_2=1) = \Pr(X_1=1) \Pr(X_2=1)$   
 $= \dots = \Pr(X_1=6, X_2=6) = 1/36$  "joint probability"

E.g., Let  $(C_1, C_2, C_3)$  be r.v.'s for the POS of 3 successive words randomly selected from a large corpus.

Possible values :  $(N, N, N), (N, N, V), \dots, (\text{Adv}, \text{Adv}, \text{Adv})$

Interdependent :  $\Pr(N, N, N) \neq (24)^3$

$$\Pr(C_1=N, C_2=N, C_3=N)$$

## Probability Axioms

1. The probability of any outcome lies between 0 and 1

$$0 \leq \Pr(X=x) \leq 1$$

where  $x$  is any possible value of r.v.  $X$

2. The probabilities of all possible alternative outcomes add up to 1. E.g.,  $\Pr(\text{FLIP}=H) + \Pr(\text{FLIP}=T) = 1$

$$\sum_{x \in \text{range}(X)} \Pr(X=x) = 1$$

Informally,

1.  $0 \leq \Pr(x) \leq 1$
2.  $\sum_x \Pr(x) = 1$

Similarly for jointly distributed r.v.'s :

1.  $0 \leq \Pr(\underline{x}) \leq 1$
2.  $\sum_{\underline{x}} \Pr(\underline{x}) = 1$

where  $\Pr(\underline{x})$  abbreviates  $\Pr(X=\underline{x})$  and  
 $\underline{X} = (X_1, \dots, X_n)$ , i.e.,  $n$  r.v.'s, and  
 $\underline{x} = (x_1, \dots, x_n)$ , i.e., an  $n$ -tuple of possible values of  $X_1, \dots, X_n$ .

## Conditional Probability

For jointly distributed r.v.'s, we may be interested in how probable a certain value is for one of them, given the value of another. Write as

$$\Pr(X=x | Y=y) \text{, or briefly, } \Pr(x|y)$$

given

E.g., Prob. that  $i^{\text{th}}$  word of a text is a noun, given that the preceding word is a determiner might be written  $\Pr(C_i = N | C_{i-1} = \text{Det})$ .

Defn  $\Pr(X=x | Y=y) = \frac{\Pr(X=x, Y=y)}{\Pr(Y=y)}$  provided that  $\Pr(Y=y) > 0$

E.g., suppose  $\Pr(C_{i-1} = \text{Det}, C_i = N) = 0.2$

i.e., when we randomly select a pair of successive words from a corpus, 20% of the time they are a Det, N pair

& suppose  $\Pr(C_{i-1} = \text{Det}) = .25$

i.e., when we randomly select a word from a corpus, 25% of the time it is a Det

$$\text{Then } \Pr(C_i = N | C_{i-1} = \text{Det}) = \frac{\Pr(C_i = N, C_{i-1} = \text{Det})}{\Pr(C_{i-1} = \text{Det})} = \frac{0.2}{0.25} = .80$$



## Chain rule

$$\Pr(a \& b) = \Pr(a) \Pr(b|a) \text{ from defn of cond'n prob.}$$

Similarly

$$\Pr(a \& b \& c) = \Pr(a) \Pr(b|a) \Pr(c|a \& b)$$

Useful esp. if we can make certain independence assumptions. Here, suppose

a, c are conditionally independent, given b:

$$\Pr(c|a \& b) = \Pr(c|b)$$

Then

$$\Pr(a \& b \& c) = \Pr(a) \Pr(b|a) \Pr(c|b)$$

We can easily generalize to n events or r.v.'s.

E.g., Suppose  $C_1, \dots, C_n$  represent the parts of speech of a stretch of n successive words randomly selected from a text corpus.

$$\Pr(C_1, \dots, C_n) = \Pr(c_1) \Pr(c_2 | c_1) \Pr(c_3 | c_1, c_2) \dots \dots \Pr(c_n | c_1, \dots, c_{n-1})$$

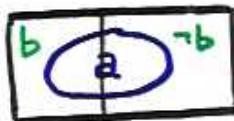
Suppose  $C_i$  is conditionally independent of all earlier parts of speech, given the immediately preceding one:

$$\Pr(c_i | c_1, \dots, c_{i-1}) = \Pr(c_i | c_{i-1})$$

$$\text{Then } \Pr(c_1, \dots, c_n) = \Pr(c_1) \prod_{i=2}^n \Pr(c_i | c_{i-1})$$

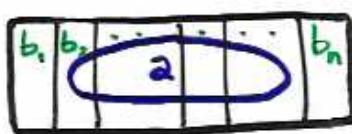
A "Markov assumption"

## "Marginal" (absolute) probabilities in terms of joint probabilities



$$\Pr(a) = \Pr(a \& b) + \Pr(a \& \neg b)$$

More generally, suppose  $b_1, \dots, b_n$  are mutually exclusive (disjoint), jointly exhaustive events (total prob. 1). Then



$$\begin{aligned}\Pr(a) &= \sum_{i=1}^n \Pr(a \& b_i) \\ &= \sum_{i=1}^n \Pr(b_i) \Pr(a|b_i)\end{aligned}$$

"Marginalization"

## Independence

Intuitively, two events  $A, B$  are independent if the fact that one occurred (or failed to occur) tells us nothing about the probability that the other occurred ; i.e.,

$$P(A|B) = P(A) \quad "B provides no evidence about A"$$

$$P(A|\neg B) = P(A)$$

$$P(B|A) = P(B)$$

$$P(B|\neg A) = P(B)$$

$$P(A \wedge B) = P(A)P(B)$$

$$P(A \wedge \neg B) = P(A)P(\neg B)$$

etc.

E.g., Flip a coin twice

Possible outcomes  $\{(H,H), (H,T), (T,H), (T,T)\}$

If all 4 prob's are  $1/4$ , then  
A, B are independent.

$\underbrace{A}_{\text{heads on 1st toss}}$        $\underbrace{B}_{\text{heads on 2nd toss}}$

Note:  
independent  
NOT disjoint

## Independent r.v.'s

E.S., For the above prob. space , let

$$X((H,H)) = X((H,T)) = 1, \quad X((T,H)) = X((T,T)) = 0$$

$$Y((H,H)) = Y((H,T)) = 1, \quad Y((T,H)) = Y((T,T)) = 0$$

i.e., "X = 1" if 1st toss is H, & 0 o.w.  
"Y = 1" if 2nd — "

## Independent r.v.'s, cont'd

So again, if  $P(X=1) = P(X=0) = P(Y=1) = P(Y=0)$ ,  
 $X, Y$  are independent r.v.'s; & we have

$$P(X=x \wedge Y=y) = P(X=x) P(Y=y) \text{ for } x, y \in \{0, 1\}$$

Also  $P(X=x | Y=y) = P(X=x)$  for  $x, y \in \{0, 1\}$   
etc.

Often abbreviations are used, such as

$P(x)$  for  $P(X=x)$ , & similarly for  $Y$ .

So, for independent  $X, Y$ ,

$$P(x, y) = P(x) P(y), \text{ etc.}$$

## Another example

Space = CS grad students

$X$  = verbal GRE score (up to 800)

$Y$  = analytical GRE score (up to 800)

$Z$  = 444 grade (up to 4.0)

These are probably nearly independent

## Conditional independence

e.g.,  $P(Z=z | X=x \wedge Y=y) = P(Z=z | X=x)$

provides no further evidence about  $Z$ ,  
beyond what  $X=x$  provides

## Bayes' Rule

$$P(A \wedge B) = P(A) P(B|A) \\ - P(B) P(A|B)$$

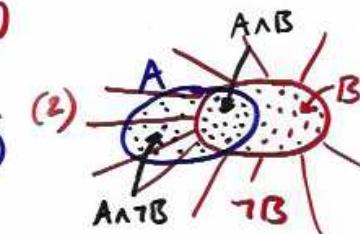
e.g.  $B = \text{cold}$   
 $A = \text{sore throat}$

$\therefore P(B|A)$  prob. of cold, given sore throat

$$= \frac{P(B) P(A|B)}{P(A)} \quad (1)$$

$$= \frac{P(B) P(A|B)}{P(A \wedge B) + P(A \wedge \neg B)} \quad (2)$$

mutually  
exclusive  
(disjoint)



$$P(B|A) = \frac{P(B) P(A|B)}{P(B) P(A|B) + P(\neg B) P(A|\neg B)} \quad (3)$$

$$P(\neg B) = \\ 1 - P(B)$$

E.g., if we know:

$$P(B) = P(\text{cold}) = .01 \quad \text{Prior prob of cold hypothesis}$$

$$P(A|B) = P(\text{sore throat} | \text{cold}) = .9$$

Likelihood of evidence, given  
the cold hypothesis

$$P(A|\neg B) = P(\text{sore throat} | \text{no cold}) = .005$$

Likelihood of evidence, given  
falsity of cold hypothesis

Then we can compute  $P(\text{hypothesis} \mid \text{new evidence})$ :

$$\begin{aligned} P(B|A) &= P(\text{cold} \mid \text{sore throat}) \\ &= \underline{\text{posterior probability of cold}} \\ &\quad \underline{\text{hypothesis}} \\ &= \frac{\text{prior}}{\cdot 01(\cdot 9)} \\ &= \frac{\cdot 01(\cdot 9)}{\cdot 01(\cdot 9) + (1 - .01)(\cdot 005)} \\ &= \underline{\cdot 645} \quad \text{NB: much bigger than prior, .01} \end{aligned}$$

### Odds-Likelihood form of Bayes' Rule

Above we needed 3 numbers to compute  $P(B|A)$ , but it really only takes 2....

By analogy with (3) :

$$P(\neg B|A) = \frac{P(\neg B)P(A|\neg B)}{P(\neg B)P(A|\neg B) + P(B)P(A|B)} \quad (3')$$

Divide (3) by (3') (Note: same denominator!):

$$\frac{P(B|A)}{P(\neg B|A)} = \frac{P(B)P(A|B)}{P(\neg B)P(A|\neg B)}$$

$$O(B|A) = O(B)L(A|B) \quad (4) \quad \begin{matrix} O & \text{"odds"} \\ L & \text{"likelihood ratio"} \end{matrix}$$

- where the "odds" in favor of an event (or proposition) A is the ratio of its probability of occurrence to its probability of nonoccurrence

$$O(B) = \frac{P(B)}{P(\neg B)} = \frac{P(B)}{1 - P(B)}$$

Similarly for conditional (or posterior) odds,

$$O(B|A) = \frac{P(B|A)}{P(\neg B|A)} = \frac{P(B|A)}{1 - P(\neg B|A)}$$

- and the likelihood ratio  $L(A|B)$  is the ratio of the likelihood of the evidence, given the hypothesis, to the likelihood of the evidence, given the denial of the hypothesis

$$L(A|B) = \frac{P(A|B)}{P(A|\neg B)}$$

### How to recover probabilities from (4) :

From  $O(B) = \frac{P(B)}{1 - P(B)}$ , it's easily shown that

$$P(B) = \frac{O(B)}{1 + O(B)}$$

Similarly

$$P(B|A) = \frac{O(B|A)}{1 + O(B|A)}$$

## Combining Evidence

E.g. might have 2 pieces of evidence for a "cold" hypothesis: sore throat, runny nose

 $H$  $E_1$  $E_2$ 

Write down Bayes rule for  $P(H|E_1)$  again:

$$\underline{P(H|E_1)} = \frac{P(H) P(E_1|H)}{P(H) P(E_1|H) + P(\neg H) P(E_1|\neg H)} \quad (5)$$

Take this posterior prob. as the prior prob. for the next piece of evidence,  $E_2$ :

$$P(H|E_1 \wedge E_2) = \frac{P(H|E_1) P(E_2|H \wedge E_1)}{P(H|E_1) P(E_2|H \wedge E_1) + P(\neg H|E_1) P(E_2|\neg H \wedge E_1)}$$

Now assume that relative to a situation where the hypothesis  $H$  is given (known to be true),  $E_1$  &  $E_2$  are independent (conditional independence):

$$P(E_2|H \wedge E_1) = P(E_2|H)$$

Similarly if  $\neg H$  is given ( $H$  is known to be false):

$$P(E_2|\neg H \wedge E_1) = P(E_2|\neg H)$$

Then (5) simplifies to:

$$P(H|E_1 \wedge E_2) = \frac{P(H|E_1) P(E_2|H)}{P(H|E_1) P(E_2|H) + P(\neg H|E_1) P(E_2|\neg H)} \quad (6)$$

So, we already know how to compute  $P(H|E_1)$  using  $P(H)$ ,  $P(E_1|H)$ , &  $P(E_1|\neg H)$

So (6) shows us how to allow for both pieces of evidence,  $E_1$  &  $E_2$ , with the additional data  $P(E_2|H)$ ,  $P(E_2|\neg H)$

E.g. Let's use  $P(\text{Cold}) = .01$ ,  $P(\text{sore throat} | \text{Cold}) = .9$ ,  $P(\text{sore throat} | \neg \text{Cold}) = .005$  as before, so we get

$$P(\text{Cold} | \text{sore throat}) = .645$$

as before.

Now assume new evidence "sniffles", where

$$P(\text{sniffles} | \text{Cold}) = .95$$

$$P(\text{sniffles} | \neg \text{Cold}) = .1$$

From (6):  $P(\text{Cold} | \text{sore throat} \wedge \text{sniffles}) =$

$$\frac{.645 (.95)}{.645 (.95) + (1 - .645) (.1)} = \boxed{.945} \quad \text{note the increase!}$$

We can get an odds-likelihood version as before : write an expression for

$$P(H|E_1 \wedge E_2)$$

analogous to (6). Call this (6'). Divide (6)/(6') :

$$\frac{P(H|E_1 \wedge E_2)}{P(\neg H|E_1 \wedge E_2)} = \frac{P(H|E_1) P(E_2|\neg H)}{P(\neg H|E_1) P(E_2|\neg H)}$$

$$O(H|E_1 \wedge E_2) = O(H|E_1) L(E_2|\neg H)$$

$$\text{But } O(H|E_1) = O(H)L(E_1|\neg H)$$

So

$$O(H|E_1 \wedge E_2) = O(H)L(E_1|\neg H)L(E_2|\neg H) \quad (7)$$

To "update" the odds in favor of  $\neg H$  as new evidence  $E_1, E_2, \dots$  comes in, we just multiply by the corresponding likelihood ratios  $L(E_1|\neg H), L(E_2|\neg H), \dots$

(but keep in mind the conditional independence assumption !)